

Stochastic Simulation

- **Idea:** probabilities \leftrightarrow samples
- Get probabilities from samples:

X	<i>count</i>
x_1	n_1
\vdots	\vdots
x_k	n_k
<i>total</i>	m

 \leftrightarrow

X	<i>probability</i>
x_1	n_1/m
\vdots	\vdots
x_k	n_k/m

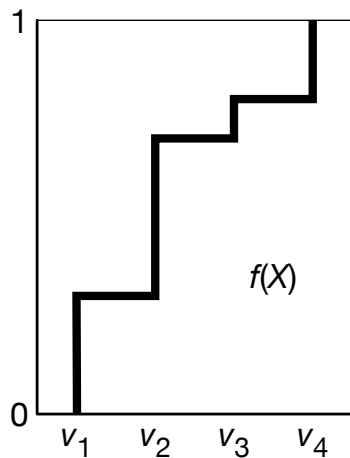
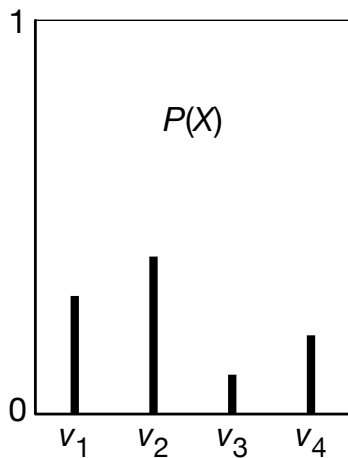
- If we could sample from a variable's (posterior) probability, we could estimate its (posterior) probability.

Generating samples from a distribution

For a variable X with a discrete domain or a (one-dimensional) real domain:

- Totally order the values of the domain of X .
- Generate the cumulative probability distribution:
 $f(x) = P(X \leq x)$.
- Select a value y uniformly in the range $[0, 1]$.
- Select the x such that $f(x) = y$.

Cumulative Distribution



Hoeffding's inequality

Theorem (Hoeffding): Suppose p is the true probability, and s is the sample average from n independent samples; then

$$P(|s - p| > \epsilon) \leq 2e^{-2n\epsilon^2}.$$

Guarantees a **probably approximately correct** estimate of probability.

If you are willing to have **an error greater than ϵ in less than δ of the cases**, solve $2e^{-2n\epsilon^2} < \delta$ for n , which gives

$$n > \frac{-\ln \frac{\delta}{2}}{2\epsilon^2}.$$

ϵ	δ	n
0.1	0.05	185
0.01	0.05	18,445
0.1	0.01	265

Forward sampling in a belief network

- Sample the variables one at a time; sample parents of X before sampling X .
- Given values for the parents of X , sample from the probability of X given its parents.

Rejection Sampling

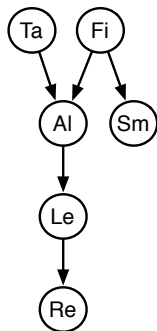
- To estimate a posterior probability given evidence $Y_1 = v_1 \wedge \dots \wedge Y_j = v_j$:
- Reject any sample that assigns Y_i to a value other than v_i .
- The non-rejected samples are distributed according to the posterior probability:

$$P(\alpha \mid \text{evidence}) \approx \frac{\sum_{\alpha \text{ is true in sample}} 1}{\sum_{\text{sample}} 1}$$

where we consider only samples consistent with evidence.

Rejection Sampling Example: $P(ta \mid sm, re)$

Observe $Sm = true, Re = true$



	Ta	Fi	Al	Sm	Le	Re	
s_1	false	true	false	true	false	false	✗
s_2	false	true	true	true	true	true	✓
s_3	true	false	true	false	—	—	✗
s_4	true	true	true	true	true	true	✓
...							
s_{1000}	false	false	false	false	—	—	✗

$$P(sm) = 0.02$$

$$P(re \mid sm) = 0.32$$

There are 1000 samples.

How many are rejected?

How many are used?

Doesn't work well when evidence is unlikely.

Importance Sampling

- Samples have weights: a real number associated with each sample that takes the evidence into account.
- Probability of a proposition is weighted average of samples:

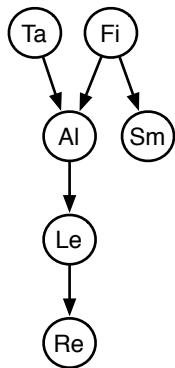
$$P(\alpha \mid \textit{evidence}) \approx \frac{\sum_{\textit{sample}: \alpha \text{ is true in sample}} \textit{weight}(\textit{sample})}{\sum_{\textit{sample}} \textit{weight}(\textit{sample})}$$

- Mix exact inference with sampling: don't sample all of the variables, but weight each sample according to $P(\textit{evidence} \mid \textit{sample})$.

Importance Sampling (Likelihood Weighting)

```
procedure likelihood_weighting( $B_n, e, H, n$ ):  
  # Approximate  $P(H \mid e)$  in belief network  $B_n$  using  $n$  samples.  
  #  $H$  has some real domain (e.g.,  $\{0, 1\}$ )  
   $mass := 0$            # mass of all samples  
   $hmass := 0$         # weighted sum of value of  $H$   
  repeat  $n$  times:  
     $weight := 1$      # weight of current sample  
    for each variable  $X_i$  in order:  
      if  $X_i = o_i$  is observed  
         $weight := weight \times P(X_i = o_i \mid parents(X_i))$   
      else assign  $X_i$  a random sample of  $P(X_i \mid parents(X_i))$   
     $mass := mass + weight$   
     $hmass := hmass + weight * (\text{value of } H \text{ in current assignment})$   
  return  $hmass/mass$ 
```

Importance Sampling Example: $P(ta \mid sm, re)$



	Ta	Fi	Al	Le	Weight
s_1	true	false	true	false	0.01×0.01
s_2	false	true	false	false	0.9×0.01
s_3	false	true	true	true	0.9×0.75
s_4	true	true	true	true	0.9×0.75
...					
s_{1000}	false	false	true	true	0.01×0.75

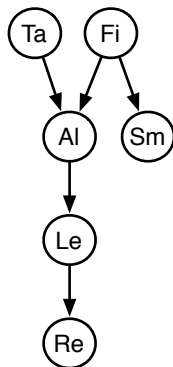
$$P(sm \mid fi) = 0.9$$

$$P(sm \mid \neg fi) = 0.01$$

$$P(re \mid le) = 0.75$$

$$P(re \mid \neg le) = 0.01$$

Importance Sampling Example: $P(le \mid sm, ta, \neg re)$



$$P(ta) = 0.02$$

$$P(fi) = 0.01$$

$$P(al \mid fi \wedge ta) = 0.5$$

$$P(al \mid fi \wedge \neg ta) = 0.99$$

$$P(al \mid \neg fi \wedge ta) = 0.85$$

$$P(al \mid \neg fi \wedge \neg ta) = 0.0001$$

$$P(sm \mid fi) = 0.9$$

$$P(sm \mid \neg fi) = 0.01$$

$$P(le \mid al) = 0.88$$

$$P(le \mid \neg al) = 0.001$$

$$P(re \mid le) = 0.75$$

$$P(re \mid \neg le) = 0.01$$

Computing Expectations & Proposal Distributions

Expected value of f with respect to distribution P :

$$\begin{aligned}\mathbb{E}_P(f) &= \sum_w f(w) * P(w) \\ &\approx \frac{1}{n} \sum_s f(s)\end{aligned}$$

s is sampled with probability P . There are n samples.
(Expectation of variable with domain $\{0, 1\}$ is its probability.)

$$\begin{aligned}\mathbb{E}_P(f) &= \sum_w f(w) * P(w)/Q(w) * Q(w) \\ &\approx \frac{1}{n} \sum_s f(s) * P(s)/Q(s)\end{aligned}$$

s is selected according to the distribution Q .

The distribution Q is called a **proposal distribution**.

$P(c) > 0$ then $Q(c) > 0$.

Try to make Q so the weights end up far from zero.

Particle Filtering

Importance sampling can be seen as:

for each particle:

for each variable:

sample / absorb evidence / update query

where **particle** is one of the samples.

Instead we could do:

for each variable:

for each particle:

sample / absorb evidence / update query

Why?

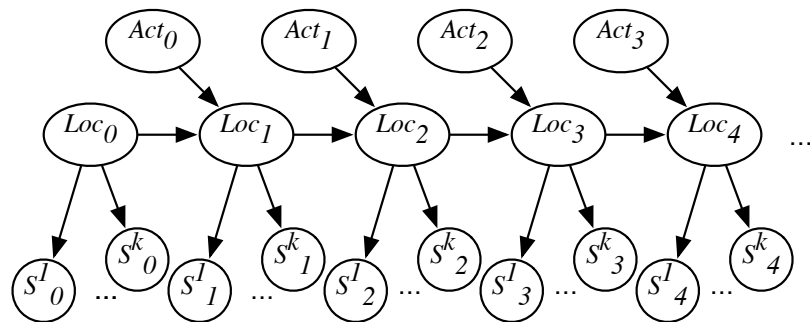
- It works with infinitely many variables (e.g., HMM)
- We can have a new operation of resampling

Particle Filtering for HMMs

- Start with random chosen particles (say 1000)
- Sample initial states in proportion to their probability.
- Repeat (as each observation arrives):
 - ▶ **Absorb evidence**: weight each particle by the probability of the evidence observation given the state of the particle.
 - ▶ **Resample**: select each particle at random, in proportion to the weight of the particle.
Some particles may be duplicated, some may be removed. All new particles have same weight.
 - ▶ **Transition**: sample the next state for each particle according to the transition probabilities.

To answer a query about the current state, use the set of particles as data.

Example: Localization

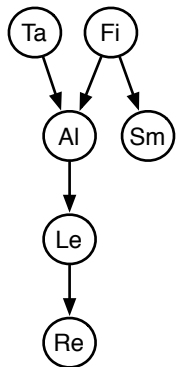


Loc consists of (x, y, θ) – position and orientation
 $k = 24$ sonar sensors (all very noisy)

To sample from a distribution P :

- Create (ergodic and aperiodic) Markov chain with P as equilibrium distribution.
Let $T(S_{i+1} | S_i)$ be the transition probability.
- Given state s , sample state s' from $T(S | s)$
- After a while, the states sampled will be distributed according to P .
- Ignore the first samples “burn-in”
— use the remaining samples.
- Samples are not independent of each other “autocorrelation”.
Sometimes use subset (e.g., 1/1000) of them “thinning”
- **Gibbs sampler**: sample each non-observed variable from the distribution of the variable given the current (or observed) value of the variables in its Markov blanket.

Gibbs Sampling Example: $P(ta \mid sm, re)$



	Ta	Fi	Al	Le
s_1	true	false	false	true
Select <i>Le</i> . Sample from $P(Le \mid \neg al \wedge re)$				
s_2	true	false	false	false
Select <i>Fi</i> . Sample from $P(Fi \mid ta \wedge \neg al \wedge sm)$				
s_3	true	true	false	false
Select <i>Al</i> . Sample from $P(Al \mid ta \wedge fi \wedge \neg le)$				
s_4	true	true	false	false
Select <i>Le</i> . Sample from $P(Le \mid \neg al \wedge re)$				
s_5	true	true	false	true
Select <i>Ta</i> . Sample from $P(Ta \mid \neg al \wedge fi)$				
s_6	true	true	false	true
...				

